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# MACHINE LEARNING AND PATTERN RECOGNITION

Fall 2004, Lecture 4:

Review of Probability and Statistics

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# Review of Probability and Statistics: Definitions

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- Random Variable  $X$ : a variable that represents a particular measurement/state of the world.
- The probability that  $X$  has value  $x$  (the result of a drawing, a sampling, or the result of a measurement) is denoted  $P(x)$ , or sometimes  $P(X = x)$ .
- The space of outcomes  $x$ , can be discrete, or continuous, possibly multidimensional.
- A discrete distribution associates a number  $0 \leq P(x) \leq 1$  to each possible outcome  $x$ , such that  $\sum_x P(x) = 1$ .
- A probability Density Function (PDF) associates a positive number  $P(x)$  to each point in the space of outcomes (can be larger than 1) such that  $\int P(x)dx = 1$ .
- The probability that  $X$  belongs to a set  $S$  is equal to  $Prob(X \in S) = \int_{x \in S} P(x)dx$ .

# Expectations

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- Expected value of a function  $f$  of a random variable  $X$  (a.k.a. the "average value"):

$$\mathcal{E}(f) = \sum_x f(x)P(x)$$

- in the continuous case:

$$\mathcal{E}(f) = \int f(x)P(x)dx$$

- Example 1, the mean of  $X$ :  $\mathcal{E}(X) = \sum_x xP(x)$

- Example 2, the variance of  $X$ :

$$\text{Var}(X) = \mathcal{E}[(X - \mathcal{E}(X))^2] = \sum_x (x - \mathcal{E}(X))^2 P(x)$$

- Example 3, the covariance of a multidimensional random variable (dimension  $N$ ):  $\text{Cov}(X) = \mathcal{E}(X.X') = \sum_x x.x'P(x)$   $x.x'$  is the outer product of  $x$  by itself:  $[x.x']_{ij} = x_i x_j$ , a symmetric  $N \times N$  matrix.

# Joint Probability

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- Two random variables  $X$  and  $Y$  (e.g.  $X$  = percentage of alcohol in the blood of a person today (continuous),  $Y = 1$  if the person is in a car crash, 0 otherwise).
- The joint probability is the function that maps an  $(x, y)$  pair to the probability that  $X = x$  and  $Y = y$  for a person.
- Dependency:  $Y$  is more likely to be 1 if  $X$  is large, and  $X$  is more likely to be large if  $Y$  is 1.
- Marginal probabilities:

$$P(x) = \sum_y P(x, y)$$

$$P(y) = \int P(x, y) dx$$

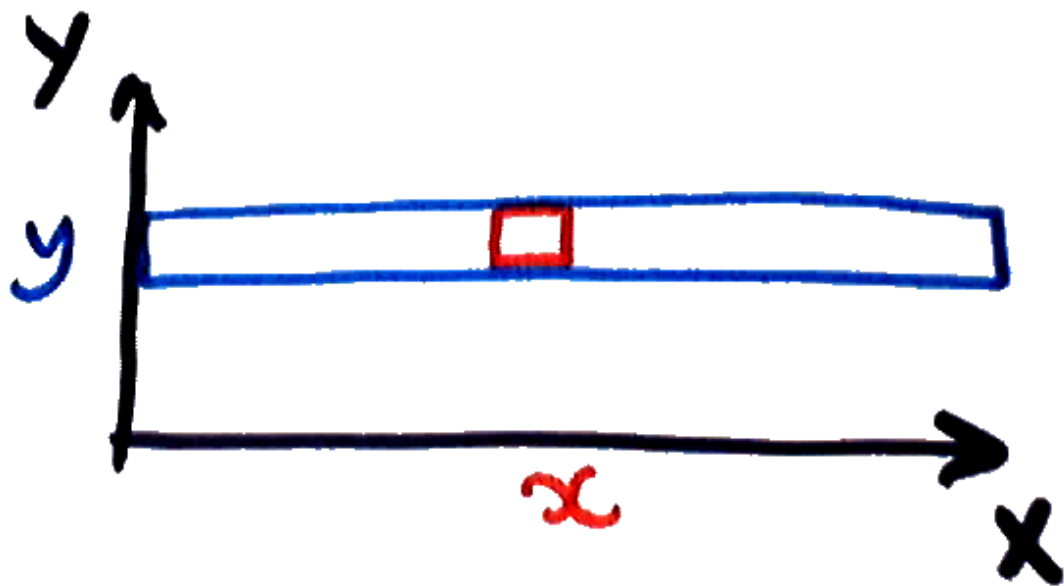
# Conditional Probability

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- Probability that someone was in a car crash knowing that the person was drunk = of all the persons who were drunk, what proportion had a car crash:

$$P(y|x) = p(x, y)/p(x)$$

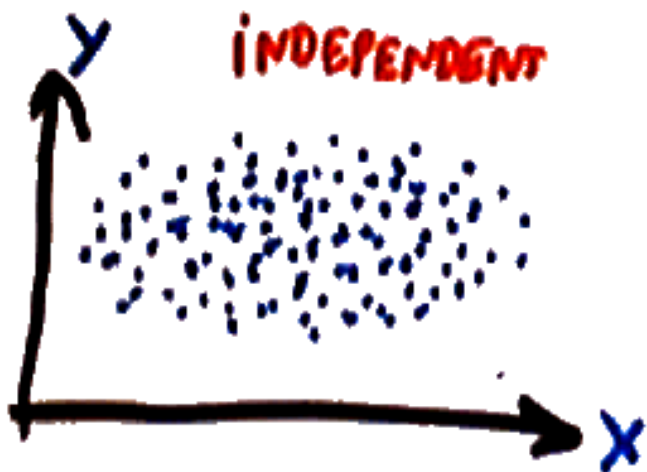
- $P(y|x)$  is read "Probability of  $y$  given  $x$ ."
- Normalization:  $\sum_y P(y|x) = 1$



# Conditional Independence

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Independence:  $X$  and  $Y$  are *independent* iff  $P(x, y) = P(x)P(y)$ , in other words  $P(x|y) = P(x)$  and  $P(y|x) = P(y)$ .



# Special Distributions: Exponential Family

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- A very general family of parameterized distributions.
- $P(x|\omega) = h(x) \exp(\omega' T(x) - A(\omega)) = \frac{1}{Z(\omega)} h(x) \exp(\omega' T(x))$
- $\omega$  the “natural” parameter
- $Z(\omega) = \exp(A(\omega))$  is the *partition function*
- $T(x)$  a *sufficient statistic*: all you need to know about  $x$  to compute its distribution with a linear combination.

# Special Distributions: Gaussian

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- For a continuous random variable:  $P(x|m, v) = \frac{1}{\sqrt{2\pi v}} \exp(-\frac{1}{2v}(x - m)^2)$
- $m$  is the mean,  $v$  is the variance.
- exponential family with
  - $w = [m/v, -1/2v]$
  - $T(x) = [x, x^2]$
  - $Z(w) = \sqrt{v} \exp(m/2v)$
  - $h(x) = 1/\sqrt{2\pi}$



# Special Distributions: Multivariate Gaussian

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- For a continuous random variable ( $X$ , and  $M$  are  $N$ -dimensional vectors,  $V$  is an  $N \times N$  matrix):

$$P(X|M, V) = |2\pi V|^{-1/2} \exp(-1/2(X - M)'V^{-1}(X - M))$$

- $|2\pi V|$  is the determinant of  $2\pi V$ .

- exponential family with

- $w = [V^{-1}M, -1/2V^{-1}]$

- $T(x) = [X, XX']$

- $Z(w) = |V|/2 \exp(1/2M'V^{-1}M)$

- $h(x) = (2\pi)^{-N/2}$

- Important facts: marginals of Gaussians are Gaussians, products of Gaussians are Gaussians, conditionals of Gaussians are Gaussians.

# Bayes' Rules

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- From the definition of conditional probabilities  $P(x, y) = P(x|y)P(y)$ .
- Therefore  $P(x, y) = P(x|y)P(y) = P(y|x)P(x)$ .
- Hence

$$P(x|y) = \frac{P(y|x)P(x)}{P(y)}$$

- Or equivalently:

$$P(x|y) = \frac{P(y|x)P(x)}{\sum_{x'} P(y|x')P(x')}$$

- This is a convenient way of reversing conditional probabilities.

# More General Forms of Bayes' Rules

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- Chain rule (any ordering works):

$$P(x, y, z) = P(x|y, z)P(y|z)P(z) = P(z|y, x)P(y|x)P(x) = \dots$$

- In general:  $P(x_1 \dots x_n) = \prod_i P(x_i | x_1 \dots x_{i-1})$  for any ordering  $1..n$ .
- Conditional Bayes inversion:

$$P(x|y, z) = \frac{P(y|x, z)P(x, z)}{P(y, z)}$$

- Chain rule and marginalization in one fell swoop (feels like a matrix-vector or matrix-matrix product):

$$P(y) = \int_x P(y|x)P(x)$$

$$P(y|z) = \int_x P(y|x)P(x|z)$$

# Probabilistic Models: Bayes Decision Theory

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A common (but according to some, flawed) way of building a classifier is to estimate the density function for each class  $P(X|C1)$  and  $P(X|C2)$ . When a new input comes in, compute the **posterior probability** of the class conditioned on the input using Bayes rule:

$$P(C1|X) = \frac{P(X|C1)P(C1)}{P(X)}$$

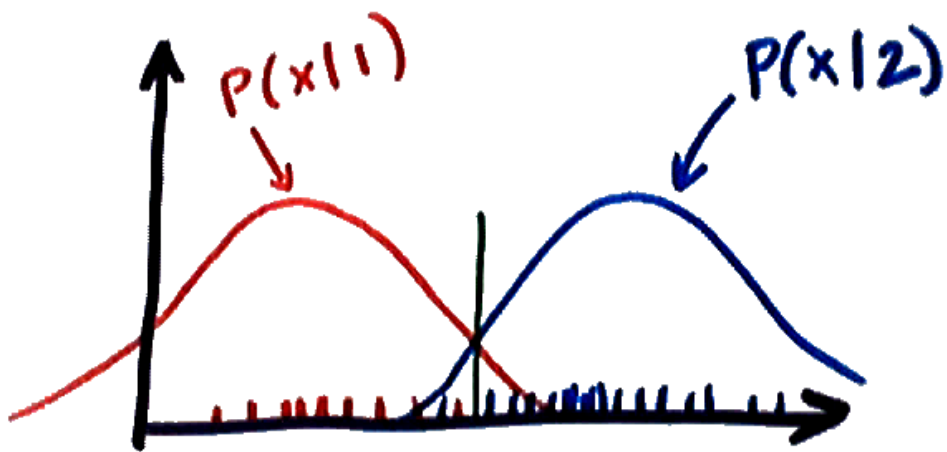
This can be rewritten as:

$$P(C1|X) = \frac{P(X|C1)P(C1)}{\sum_C P(X|C)P(C)}$$

The same can be done for class  $C2$ . Then, pick the class that has the largest posterior probability for the given  $X$ .

# Minimum Bayes Error Rate

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The area of the intersection between the two curves (assuming those curves are the real ones, not just estimates) is the **Minimum Bayes Error Rate**. Inputs that fall into that region are always classified wrong by the Bayes decision rule.

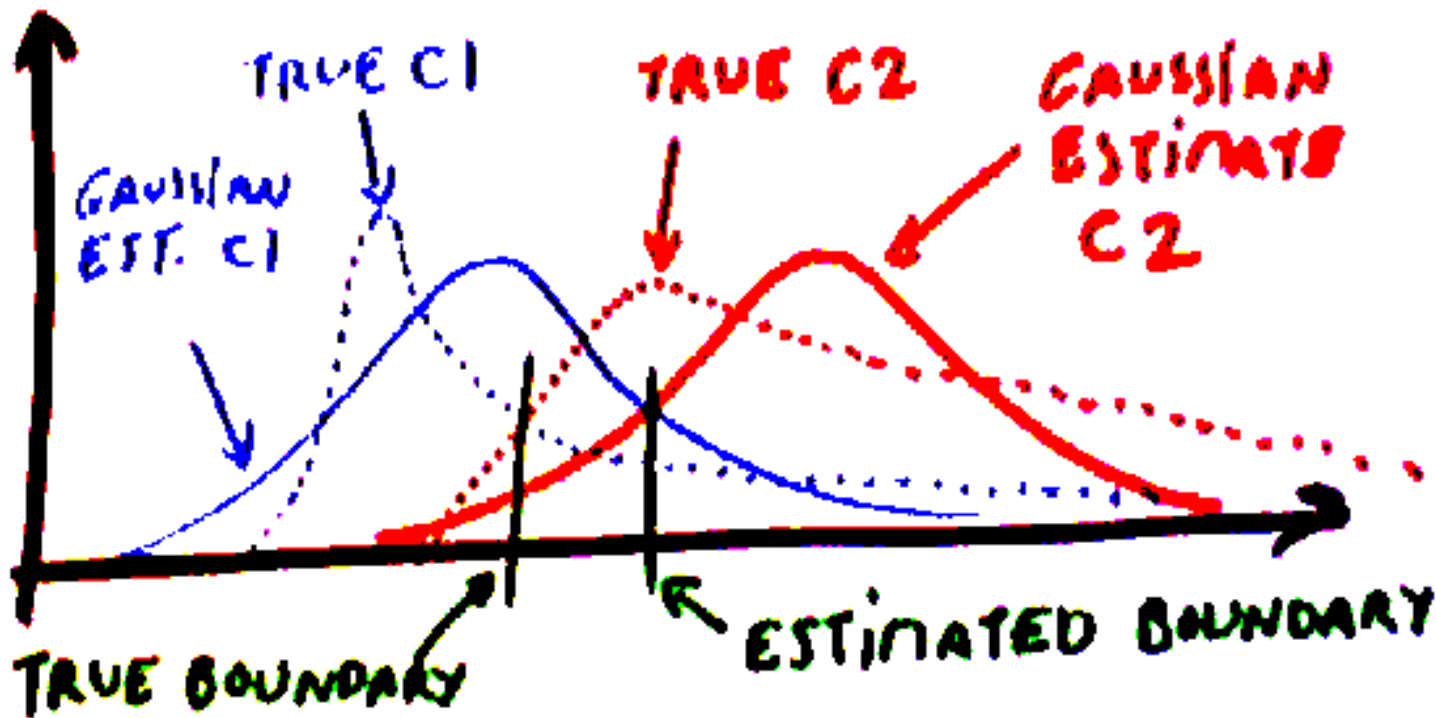
**CAUTION:** in practice we **never** know the “real” distributions, so we can never really compute the Bayes error rate, except in datasets that we cook up artificially by sampling from known distributions.

In real life *there is no such thing as “the distribution from which the data is sampled”*, we are just given a finite number of samples, period.

Assuming that our samples are drawn independently from some distribution is a convenient (sometimes necessary) hypothesis, but we must keep in mind that it's *wrong*.

# Generative Classifiers, Flawed?

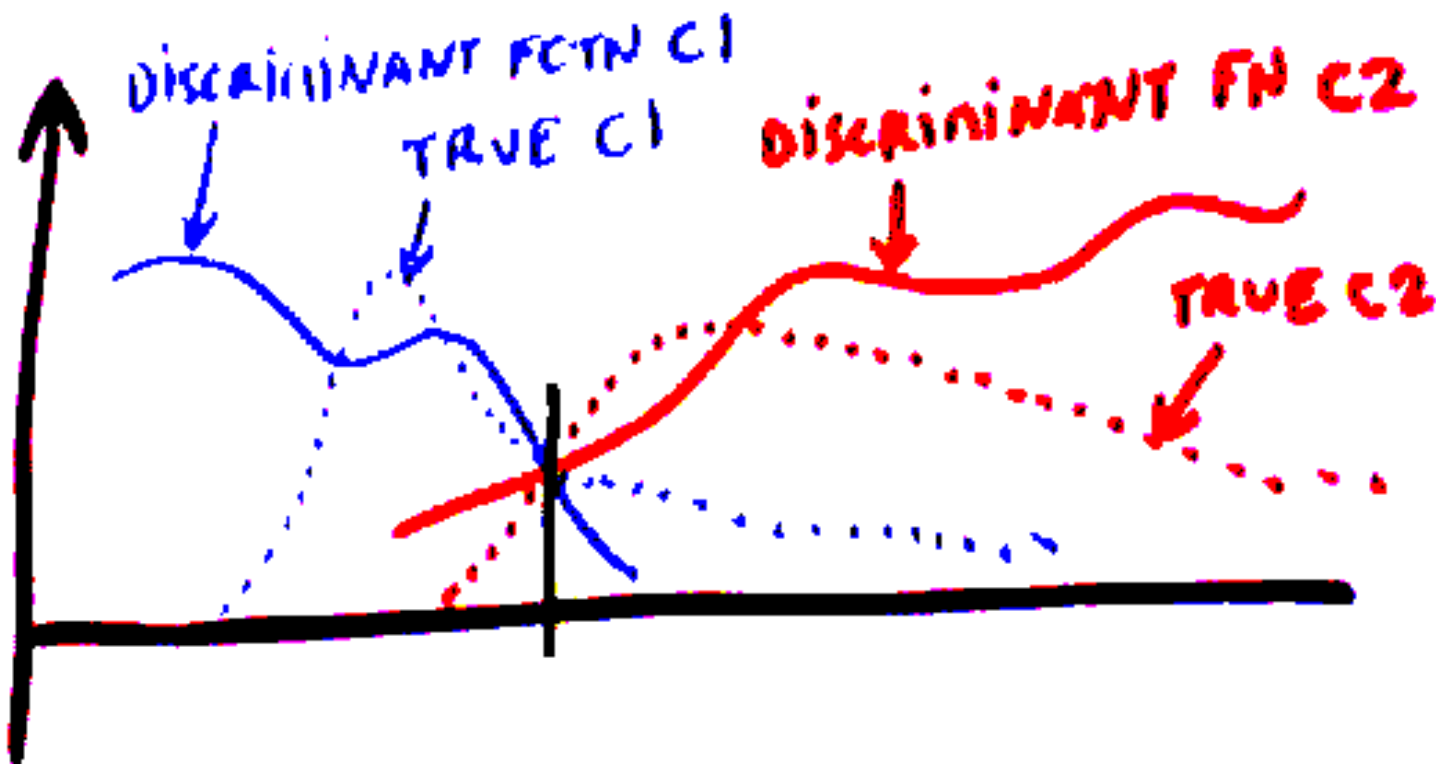
A common criticism of Bayesian classifiers and other **generative** models is that they require us to solve a much more complicated problem than we have to. We are asked to solve several density estimation problems over the whole space just to come up with a decision boundary.



# Discriminative Classifiers

Discriminative classifiers (such as the Perceptron) do not attempt to estimate the class densities, but simply try to find a suitable boundary (or simply try to estimate the class posterior probabilities without going through the class densities).

This is a considerably easier problem than estimating densities over the whole space.



# Naive Bayes Classifier

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- The Naive Bayes classifiers is a very simple (but way suboptimal) linear classifier. It assumes independence of the input variables.
- Simple setting: two class classification problem
- Probability that  $X$  belong to class  $C1$ :

$$P(C1|X) = P(X|C1)P(C1)/P(X)$$

Where  $P(X)$  is simply  $P(X|C1)P(C1) + P(X|C2)P(C2)$ .

- Let's assume that the input variables  $x_i$  are independent, we can factorize  $P(X|C1)$  as a product  $\prod_i P(x_i|C1)$ :

$$P(C1|X) = \frac{\prod_i P(x_i|C1)P(C1)}{P(X)}$$



# Naive Bayes Classifier

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- Estimating the terms  $P(x_i|C1) = P(x_i, C1)/P(C1)$  is simply performed by counting of how many times the  $i$ -th input variable takes the value  $x_i$  when the sample category is  $C1$ , and dividing by the number of samples of class  $C1$ .
- To classify, we can drop the constant term  $P(X)$  (which does not change from class to class). Taking logs we can write:

$$\log P(C1|X) = \log P(C1) + \sum_i \log[P(x_i|C1)]$$

If the variables  $x_i$  are binary (1 or 0) we can write this as

$$\log P(C1|X) = \log P(C1) + \sum_i (1-x_i) \log[P(x_i = 0|C1)] + x_i \log[P(x_i = 1|C1)]$$

# Naive Bayes Classifier

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- regrouping the terms:

$$\log P(C1|X) = \log P(C1) + \sum_i \log[P(x_i = 0|C1)] +$$

$$\sum_i (\log[P(x_i = 1|C1)] - \log[P(x_i = 0|C1)])x_i$$

This is just like a linear classifier of the form  $W_0 + W'X$  with funny weights and biases. Naive Bayes classifiers rarely work well compared to discriminative linear classifiers.

# Estimating Probabilities

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- Estimating probabilities cannot be performed without a **model**, a set of independence hypotheses, and a well defined set of measurements.
- Since those choices are somewhat arbitrary, there is no such thing as “The Probability” of a real event, there are only estimates conditioned upon arbitrary assumptions.
- Example: I toss a fair coin, here is the result:  
110111001110111101000000110100...
- Now, predict the next toss.
- Method 0 [charming naïveté]: you told me it was a fair coin, so 0 and 1 are equiprobable.
- Method 1 [independent draws]: I assume that the draws are independent (the next bit does not directly depend upon the previous bits). I Just compute the empirical ratio of 1 and 0 and predict accordingly.

# Estimating Probabilities

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- 110111001110111101000000110100...
- Method 2 [extra measurements]: If I use my secret super-duper measurement device, I can get a glimpse of the state of the universe within cubic kilometer around you (including your brain). With that, I can predict which side the coin will fall on with quasi-certainty (except for quantum interactions with the rest of the universe). Each bit now depends on  $10^{100}$  known bits (and an even larger number of unknown but largely irrelevant bit) through a horribly complicated function.
- Method 3 [internal structure/dependencies]: I know you cooked up this example. Those bits would not have something to do with the decimals of  $\pi$  by any chance?

Depending on your hypotheses and assumptions, your probability estimate may be very different from mine.

# Probabilistic Linear Classification: Logistic Regression

- We want to classify vectors into two classes  $C1$  and  $C2$ .
- We assume that the quantity  $\log \frac{P(C1|X,W)}{P(C2|X,W)}$  is parameterized as a linear combination of the inputs ( $W$  is the parameter vector):

$$\log \frac{P(C1|X, W)}{P(C2|X, W)} = W' X$$

- since we only have two classes, we can write  $P(C2|X, W) = 1 - P(C1|W, X)$
- hence

$$\frac{P(C1|X, W)}{1 - P(C1|X, W)} = \exp(W' X)$$

- solving for  $P(C1|X, W)$ , we get:

$$P(C1|X, W) = \sigma(-W' X) = \frac{1}{1 + \exp(-W' X)}$$

$\sigma$  is called the logistic function.

# Estimating a Logistic Regression

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- How do we compute the  $W$  that best approximates the desired distribution of  $P(C|X)$ ?
- We measure the “distance” between the desired distribution (which is given by the samples) and the proposed distribution.
- A good dissimilarity measure between two discrete distributions  $P$  and  $Q$  is the **Kullback-Leibler Divergence**:

$$KL(Q, P) = - \sum_x Q(x) \log(P(x)/Q(x))$$

- in our case:

$$L(W) = - \sum_i y^i \log(P(C1|X^i)) + (1 - y^i) \log(1 - P(C1|X^i))$$

where  $y^i$  is 1 if sampl  $X^i$  is of class 1, and 0 if it is of class 2.

# Estimating a Logistic Regression

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- Logistic regression objective function:

$$L(W) = - \sum_i y^i \log(P(C1|X^i)) + (1 - y^i) \log(1 - P(C1|X^i))$$

where  $y^i$  is 1 if sampl  $X^i$  is of class 1, and 0 if it is of class 2.

- We can minimize  $L(W)$  by gradient descent:

$$W \leftarrow W - \eta \frac{\partial L(W)}{\partial W}$$

with

$$\frac{\partial L(W)}{\partial W} = \sum_i (y^i - \sigma(W' X^i)) X^i$$

- This looks a lot like the Perceptron learning rule