

Lecture #2

- First and Second Moment Methods
- Probabilistic Method
- Random Graphs (Erdős-Renyi Random Graph Model).

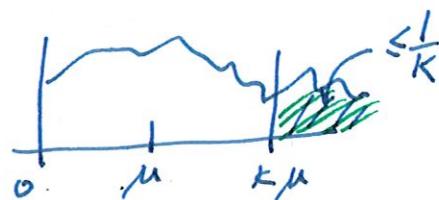
Specific Ideas

① Markov's Inequality

"first moment"
mean
↓
if X is a random variable (and non-negative) with $E[X] = \mu$

then $\Pr(X \geq k\mu) \leq \frac{1}{k}$. $\forall k \geq 0$

$$\Leftrightarrow \Pr(X \geq \lambda) \leq \frac{E(X)}{\lambda} \quad \forall \lambda \geq 0$$



"At most half the people can be at least twice the average".

② Chebyshov's Inequality

"concentration of measure"
 X is r.v. (not necc. non-neg anymore) with mean $E[X] = \mu$

And variance $E[(X-\mu)^2] = \sigma^2$

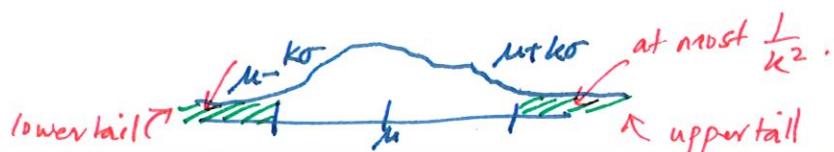
σ = standard deviation

σ^2 = variance

"second moment"

then

$$\Pr(|X-\mu| \geq k\sigma) \leq \frac{1}{k^2}.$$



Probabilistic method

Want to show that some object exists.

Set up a random experiment such that

$$\Pr(\text{get object at the end}) > 0$$

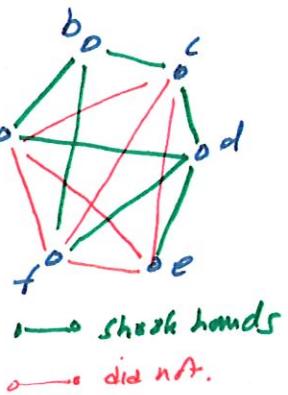
\Rightarrow the object must exist

Thm: At any party you invite 6 people. Some handshakes happen.

Then either $\exists 3$ people who all shook hands with each other
or $\exists 3$ — none of whom shake hands with any other.

Rephrase I: \exists a green Δ or a red Δ .

in a 2-coloring of the complete graph K_6 .



PF: s.p.s vrtx \$a\$ has ^{green} degree ≥ 3 , say \$b, c, d\$.

if a green edge between any pair \$\{b, c\}, \{c, d\}, \{b, d\} \Rightarrow\$ green Δ .

else no ^{green} edge $\Rightarrow \{b, c, d\}$ is a red Δ .

case of \$a\$ having red degree ≥ 3 same.



Drop all red edges. (No edge = red edge, edge = green edge)

Rephrase II:

Every graph on 6 vertices has either a clique of size 3
or an independent set of size 3.

Ramsey-type theorem.

Thm 2: Every graph on \underline{n} vertices has either

- a clique of size s or
- an independent set of size t

if $2^{s+t} \leq n+1$

PF: by induction on $s+t$.

if $s=1 \text{ or } t=1$, trivial

$s=2 \text{ or } t=2$, easy : if a graph has ~~an~~ edge (then has $K_2 = K_s$)
 } or else is independent set.

similar for $t=2$

Now suppose true for $s+t = k-1$.

to prove for $s+t = k$, consider graph on $n \geq 2^{k-1}$ vertices.

Pick v . either has $\geq \frac{n-1}{2} = 2^{k-1}-1$ ~~neigh~~ neighbors.

\Rightarrow recurse. has either clique of size $s-1$ \leftarrow combine with v to
 or I.S of size t \leftarrow get ~~K_s clique~~

else is not adjacent to $\geq 2^{k-1}-1$ vertices.

Find clique of size s or I.S of size $t-1$. \leftarrow extend with v to
 get I.S of size t .



Corollary: Every graph on n vertices has either

clique of size $\frac{1}{2} \log_2 n$

or I.S of size $\frac{1}{2} \log_2 n$

Can we do better?

: No! :

(up to constants at least)

How to prove

Need magical graph that has no large clique $\xrightarrow{\text{of size } 2\log n}$
 and no large IS $\xrightarrow{\text{of size } 2\log n}$ too!

Use the Probabilistic Method!

Consider the random graph $G(n, p)$

$$n \in \mathbb{Z}_{\geq 1}$$

$$p \in [0, 1].$$

n vertices

each of the $\binom{n}{2}$ possible edges present independently with prob. p .

_____ x _____

Thm 4: $G \sim G(n, \frac{1}{2})$ has no indep. set or clique of size $2\log n + 1$. $\xrightarrow{\text{with high prob.}}$
 "G sampled from the distribution described above"

Pf: $\forall S \subseteq V$ define indicator random variable

$$X_S = \begin{cases} 1 & \text{if } S \text{ clique in } G \\ 0 & \text{ow.} \end{cases} \Leftrightarrow X_S = \mathbf{1}_{S \text{ is a clique in } G}$$

$$\textcircled{A} \quad \mathbb{E}[X_S] = \left(\frac{1}{2}\right)^{\#\text{edges in } S} = \left(\frac{1}{2}\right)^{\binom{|S|}{2}} \quad (\text{independence of each edge})$$

$$\text{if } |S| = 2\log n + 1 = \$ \Rightarrow \mathbb{E} X_S = \left(\frac{1}{2}\right)^{\frac{S(S-1)}{2}}$$

$$\textcircled{B} \quad \mathbb{E}[\#\text{cliques}] = \binom{n}{S} \left(\frac{1}{2}\right)^{\frac{S(S-1)}{2}} \quad (\text{linearity of expectation})$$

$$\mathbb{E}\left[\sum_{S \subseteq V}^{\#\text{cliques}} X_S\right] = \sum_{S \subseteq V}^{\#\text{cliques}} \mathbb{E} X_S$$

$$|S|=s \quad |S|=s$$

Let $X = \# \text{ of cliques of size } s = \sum_{|S|=s} X_S$

$$\begin{aligned}
 \text{Some algebra: } \mathbb{E}X &= \binom{n}{s} \cdot 2^{-\frac{s(s-1)}{2}} \\
 &\leq \frac{n^s}{s!} \cdot \left(2^{-\frac{(s-1)s}{2}}\right)^s \\
 &= \frac{1}{s!} = o(1) \quad \text{↑ goes to 0 as } n \rightarrow \infty \\
 &\quad \left(\text{even for } s=3, \leq \frac{1}{3}\right)
 \end{aligned}$$

but $\binom{n}{s} \leq \frac{n^s}{s!}$
 $\Rightarrow 2^{-\frac{(s-1)s}{2}} = 2^{-\log_2 n} = \frac{1}{n}$

so $\mathbb{E}(X) \leq o(1)$

$$\Pr(X \neq 0) = \Pr(X \geq 1) \stackrel{\text{Markov}}{\leq} \mathbb{E}(X) \leq o(1).$$

$\Pr(G \text{ has a } t\text{-clique})$

Similarly: if $Y_s = \mathbf{1}(S \text{ is indep set})$ and $Y = \sum_{|S|=s} Y_S$

then $\Pr(Y \neq 0) \leq o(1)$.

Final Step: Union Bound.

$$\Pr(A \cup B) \leq \Pr(A) + \Pr(B)$$

$\Pr(G \text{ has a clique of size } s \text{ or an indep set of size } s) = \Pr(X \neq 0 \text{ or } Y \neq 0)$

$$\leq \Pr(X \neq 0) + \Pr(Y \neq 0)$$

$$= o(1).$$

in fact almost all graphs of this size have this property!

Corollary: \exists a graph on n nodes with no s -clique or s -IS when $s = 2\log n + 1$.

Recap: run some random experiment.

define: for each forbidden ^{structure} object S , $X_S = \mathbb{1}(S \text{ exists})$

$$X := \sum_S X_S = \# \text{ forbidden objects}$$

Show: $\mathbb{E}X < 1$

✓ Markov (first moment)

Use Markov: $\Pr(X \neq 0) = \Pr(X \geq 1) \leq \mathbb{E}X < 1$

$\Rightarrow \Pr(X=0) > 0 \Rightarrow \exists \text{ solution with no forbidden structure.}$

$$\xrightarrow{\hspace{1cm}} X \xleftarrow{\hspace{1cm}}$$

Balls & Bins

n balls

n bins

Each ball thrown into a uniformly at random bin
"u.a.r"

Fact: $E[\text{load of bin } \#1] = (\#\text{balls}) \cdot \left(\frac{1}{\#\text{bins}}\right) = \frac{n}{n} = 1$

↑ linearity of expectations

But what is maximum load ~~of~~ (over all bins)

Theorem: with ~~n~~ balls, maximum load = $O\left(\frac{\log n}{\log \log n}\right)$ whp.

Pf: Set of $s = \frac{8 \log n}{\log \log n}$ balls

$$E[X_s] = \sum_{|S|=s} E(X_S) = \binom{n}{s} \cdot \left(\frac{1}{n}\right)^s \cdot n$$

of S .
all balls into n bins

$$\begin{cases} X_S = \mathbb{1}(S \text{ goes into the same bin}) \\ X = \sum S \end{cases}$$

$$= n \cdot \binom{n}{s} \cdot \left(\frac{1}{n^s}\right)$$

$$\leq n \cdot \frac{n^s}{s!} \cdot \frac{1}{n^s}$$

$$\binom{n}{s} \leq \frac{n^s}{s!}$$

But $s! = 1 \cdot 2 \cdot 3 \dots s$
 $\geq (1 \cdot s) (2 \cdot (s-1)) (3 \cdot (s-3)) \dots \geq s^{\frac{s}{2}}$
 $s_{\frac{s}{2}} \text{ terms}$

$$\leq \frac{n}{s^{\frac{s}{2}}} \quad \text{But } s = \frac{8 \log n}{10 \log n} \text{ say.} \Rightarrow s^{\frac{s}{2}} \geq n^2 \text{ (algebra)}$$

~~Please try and if you have~~

$$\leq \frac{1}{n}$$

$\Rightarrow \Pr(\exists s \text{ of sizes that falls in same bin})$

$$= \Pr[X \geq 1] \leq E[X] \leq \frac{1}{n}.$$

$\Rightarrow \Pr(\max \text{load} \leq s) \text{ w.p. } 1 - \frac{1}{n}$



$$\begin{aligned} s^{\frac{s}{2}} &\geq (\sqrt{10 \log n})^{\frac{4 \log n}{10 \log n}} \\ &= (2^{\frac{4 \log n}{10 \log n}})^{\frac{4 \log n}{10 \log n}} \\ &= 2^{2 \log n} = n^2 \end{aligned}$$

— X —

Good: used 1st Moment method to show.

(a) $G(n, p=1/2)$ does not have a clique of size $2 \log n + 1$ w.h.p.

(b) no bin has $\geq \frac{8 \log n}{10 \log n}$ balls w.h.p.

Are these results tight? Yes!

How to prove? SECOND MOMENT METHOD!

Second Moment Method

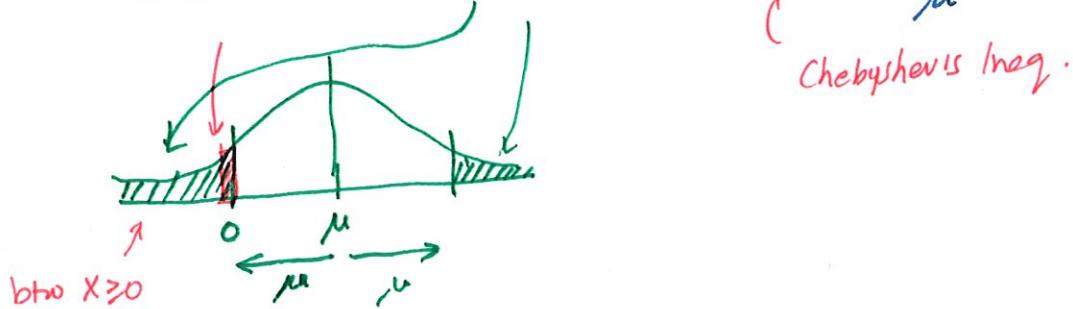
Collection ℓ of good items. We generate a random graph say.

$X_S = \mathbb{1}_{\{ \text{random graph contains item } S \text{ from } \ell \}}$

$X = \sum_{S \in \ell} X_S = \# \text{good items we contain in } G$

Want to show $P(X \neq 0) \xrightarrow{\text{with high probability}} 1-o(1)$
 i.e. $P(X=0) \text{ with tiny probability } o(1)$.

$$P_{\mathcal{X}}(X=0) \leq P_{\mathcal{X}}(|X-\mu| \geq \mu) \leq \frac{\text{Variance}(X)}{\mu^2} \quad (\mu = EX)$$



$$\Pr(X \leq 0) = \Pr(X=0)$$

How does $\text{variance}(X)$ relate to $(E(X))^2$

$$\begin{aligned} & E(X-\mu)^2 \\ &= E(X^2) - E(X)^2 \end{aligned}$$

to show that $\Pr(X=0) = o(1)$

want that $E(X^2) = (1+o(1))(E(X))^2$

want $\text{Var}(X) = o(1), E(X)$

Recall: $E(X^2) \geq E(X)^2$ by Jensen's Inequality.

~~at least~~:

Let's do it for balls and bins

Cliques in
(Random Graphs later or in HWs).

Thm: The max load for n -balls n -bins is $\Omega\left(\frac{\log n}{\log \log n}\right)$ whp.

OK: Again ~~$X_i = \mathbb{1}(\text{all balls in } S \text{ fall in bin } i)$~~

$$s = \frac{\log n}{3 \log \log n}$$

This time $X_i = \mathbb{1}(\text{some set } A \ni s \text{ balls falls into } i)$. for bin i

$$E(X_i) \geq \Pr(\text{Bin } i \text{ has exactly } s \text{ balls})$$

$$= \sum_{|S|=s} \Pr(\text{Bin } i \text{ gets exactly the balls in } S)$$

$$= \binom{n}{s} \cdot \left(\frac{1}{n}\right)^s \cdot \left(1 - \frac{1}{n}\right)^{n-s}$$

$$\geq \left(\frac{n}{s}\right)^s \cdot \frac{1}{n^s} \cdot \frac{1}{4} \geq \frac{1}{4s^s} \geq \frac{1}{n^{s/3}}$$

$$\cdot \binom{n}{s} \geq \left(\frac{n}{s}\right)^s$$

$$\cdot \left(1 - \frac{1}{n}\right)^{n-s} \geq \left(1 - \frac{1}{n}\right)^n \geq \frac{1}{4}$$

for our choice of s .

def $X = \sum_i X_i$
 $= \#\text{bins with } \geq s \text{ balls.}$

$$\Rightarrow E[\#\text{bins with at least } s \text{ balls}] \geq n \cdot \frac{1}{n^{s/3}} = n^{2/3}$$

$$E(X) = \sum_i E(X_i)$$

Linearity of Expectation

Great: So we expect $n^{2/3}$ bins to have many balls. in Expectation

But can we infer that some bin will have many balls? whp

Not yet!

with high prob
ie. w.p. $1 - o(1)$?

(E.g. maybe w.p. $\frac{1}{2}$, each bin has 1 ball)

w.p. $\frac{1}{2}$, $2n^{2/3}$ bins have lots of balls)

not possible, as we will see, but we need more work...

To use the second moment idea, we say:-

$$P_r(X=0) \leq P_r(|X - EX| \geq EX) \stackrel{\text{chebyshev}}{\leq} \frac{Var(X)}{(EX)^2} \leq \frac{Var(X)}{n^{4/3}}.$$

What is $Var(X)$? Remember $X = \sum_{i=1}^n X_i$

$$Var(X) = \sum_i Var(X_i) + \sum_{i \neq j} Covariance(X_i, X_j)$$

Fact 1 In this experiment

$$Cov(X_i, X_j) \leq 0$$

" X_i and X_j are negatively correlated: if $X_i=1$

then X_j is less likely to be 1". (Proof later)

$$\begin{aligned} Cov(X_i, X_j) &= E[(X_i - \mu_i)(X_j - \mu_j)] \\ Var(X_i) &= Cov(X_i, X_i) \\ &= E[(X_i - \mu_i)^2] \end{aligned}$$

So: $Var(X) \leq \sum_i Var(X_i) \stackrel{\text{all bins identically distributed.}}{=} n \cdot Var(X_i)$

Fact 2: Since $X_i \in \{0, 1\}$, $Var(X_i) \leq E(X_i)$

$$\text{Pf: } Var(X_i) = E(X_i^2) - E(X_i)^2 \leq E(X_i^2) = E(X_i) \quad \text{since } \begin{cases} \mu^2 = 1 \\ \sigma^2 = 0. \end{cases}$$

$$P_r(X=0)$$

$$\Rightarrow \cancel{\frac{Var(X)}{(EX)^2}} \leq \frac{Var(X)}{(EX)^2} \leq \frac{E(X)}{(EX)^2} = \frac{1}{EX} \leq \frac{1}{n^{4/3}}$$

$$\Rightarrow \text{with prob } 1 - \frac{1}{n^{2/3}}, \text{ some bin has } \geq \frac{\log n}{3 \log \log n} \text{ balls} \quad \text{:(}$$

Great: Similarly you can show (HW) that

whp, the graph $G_{n,1/2}$ contains a clique of size $(2-\varepsilon)\log n$.

[Diff: cannot just use negative correlation, need to do some work]
(not true) to control $E[X_5 X_7]$

$$\longrightarrow X \longrightarrow$$

Wrap up: actually, tighter analysis can show that

#balls in most loaded bin is $(1 \pm o(1)) \cdot \frac{\ln n}{\ln \ln n}$ with high prob $1-o(1)$.

[see notes on webpage from MU book]

$$\longrightarrow X \longrightarrow$$

Cool: today saw

(1) Markov's Ineq : $\Pr(X \geq \lambda) \leq \frac{E(X)}{\lambda}$ if $X \geq 0$.

(2) Chebyshov : $\Pr(|X-\mu| \geq \lambda) \leq \frac{\text{Var}(X)}{\lambda^2}$ for any random var.

(3) First moment method : if $E(X) = o(1)$
then $\Pr(X=0) \geq 1 - E(X) = 1 - o(1)$.
↑ whp
zero \Rightarrow bad things happen.

(4) Second moment method : $X = \# \text{good things}$

if $\frac{\text{Var}(X)}{E(X)^2} = o(1)$ then

$\Pr(X=0) \leq \frac{\text{Var}(X)}{(Ex)^2} = o(1) \Rightarrow \Pr(X \geq 1) \geq 1 - o(1).$
↑ whp
at least one good thing happens

