

LECTURE 8:  
LATENT VARIABLE MODELS

Sam Roweis

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PARTIALLY UNOBSERVED (MISSING) VARIABLES

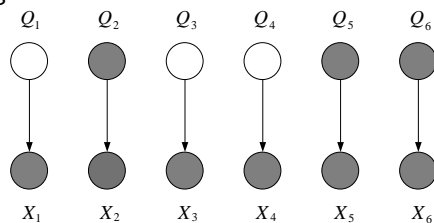
- If variables are occasionally unobserved they are *missing data*.  
e.g. undefined inputs, missing class labels, erroneous target values
- In this case, we can still model the joint distribution, but we define a new cost function in which we *sum out* or *marginalize* the missing values at training or test time:

$$\begin{aligned} \ell(\theta; \mathcal{D}) &= \sum_{\text{complete}} \log p(\mathbf{x}^c, \mathbf{y}^c | \theta) + \sum_{\text{missing}} \log p(\mathbf{x}^m | \theta) \\ &= \sum_{\text{complete}} \log p(\mathbf{x}^c, \mathbf{y}^c | \theta) + \sum_{\text{missing}} \log \sum_{\mathbf{y}} p(\mathbf{x}^m, \mathbf{y} | \theta) \end{aligned}$$

[Recall that  $p(x) = \sum_q p(x, q)$ .]

UNOBSERVED VARIABLES

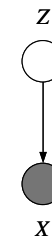
- Certain variables  $Q$  in our models may be *unobserved*, either some of the time or always, either at training time or at test time.



Graphically, we will use shading to indicate observation.

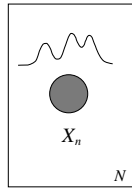
LATENT VARIABLES

- What to do when a variable  $z$  is *always* unobserved?  
Depends on where it appears in our model. If we never condition on it when computing the probability of the variables we *do* observe, then we can just forget about it and integrate it out.  
e.g. given  $\mathbf{y}, \mathbf{x}$  fit the model  $p(\mathbf{z}, \mathbf{y} | \mathbf{x}) = p(\mathbf{z} | \mathbf{y})p(\mathbf{y} | \mathbf{x}, \mathbf{w})p(\mathbf{w})$ .  
(In other words if it is a leaf node.)
- But if  $z$  is conditioned on, we need to model it:  
e.g. given  $\mathbf{y}, \mathbf{x}$  fit the model  $p(\mathbf{y} | \mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{y} | \mathbf{x}, \mathbf{z})p(\mathbf{z})$

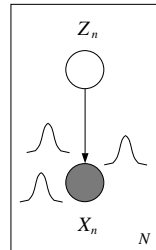


## WHERE DO LATENT VARIABLES COME FROM?

- Latent variables may appear naturally, from the structure of the problem, because something wasn't measured, because of faulty sensors, occlusion, privacy, etc.
- But also, we may want to *intentionally* introduce latent variables to model complex dependencies between variables without looking at the dependencies between them directly.  
This can actually simplify the model (e.g. mixtures).



(a)



(b)

## WHY IS LEARNING HARDER?

- In fully observed iid settings, the probability model is a product thus the log likelihood is a sum where terms decouple.  
(At least for directed models.)

$$\begin{aligned}\ell(\theta; \mathcal{D}) &= \log p(\mathbf{x}, \mathbf{z} | \theta) \\ &= \log p(\mathbf{z} | \theta_z) + \log p(\mathbf{x} | \mathbf{z}, \theta_x)\end{aligned}$$

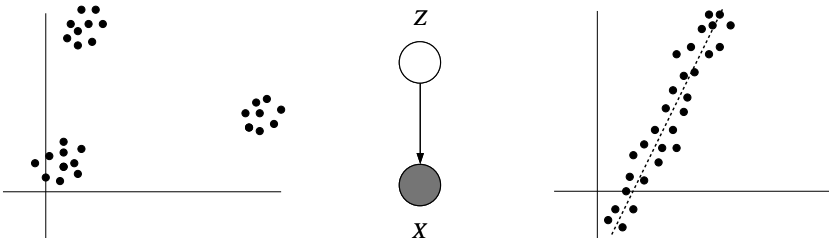
- With latent variables, the probability already contains a sum, so the log likelihood has all parameters coupled together via  $\log \sum()$ :

$$\begin{aligned}\ell(\theta; \mathcal{D}) &= \log \sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z} | \theta) \\ &= \log \sum_{\mathbf{z}} p(\mathbf{z} | \theta_z) p(\mathbf{x} | \mathbf{z}, \theta_x)\end{aligned}$$

(Just as with the partition function in undirected models.)

## CLUSTERING VS. CLASSIFICATION LATENT FACTOR MODELS VS. REGRESSION

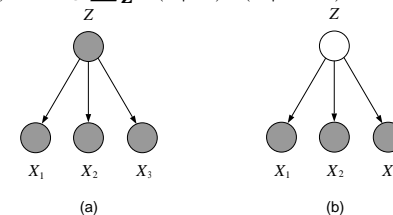
- You can think of clustering as the problem of classification with missing class labels.



- You can think of factor models (such as factor analysis, PCA, ICA, etc.) as linear or nonlinear regression with missing inputs.

## LEARNING WITH LATENT VARIABLES

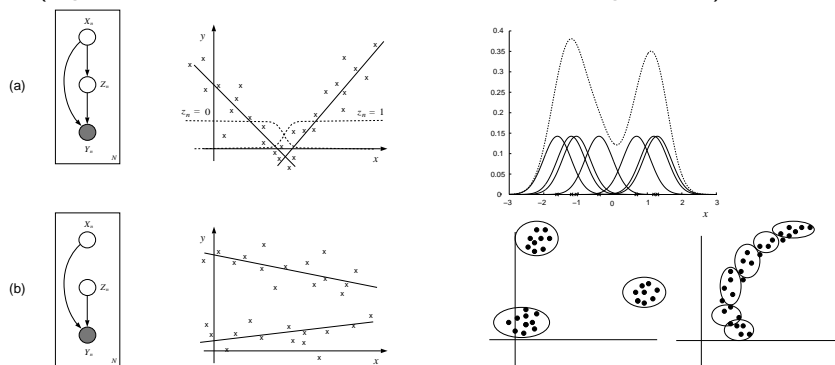
- Likelihood  $\ell(\theta; \mathcal{D}) = \log \sum_{\mathbf{z}} p(\mathbf{z} | \theta_z) p(\mathbf{x} | \mathbf{z}, \theta_x)$  couples parameters:



- We can treat this as a black box probability function and just try to optimize the likelihood as a function of  $\theta$  (e.g. gradient descent). However, sometimes taking advantage of the latent variable structure can make parameter estimation easier.
- Good news: soon we will see the *EM algorithm* which allows us to treat learning with latent variables using fully observed tools.
- Basic trick: guess the values you don't know.  
Basic math: use convexity to lower bound the likelihood.

## MIXTURE MODELS

- Most basic latent variable model with a single discrete node  $z$ .
- Allows different submodels (experts) to contribute to the (conditional) density model in different parts of the space.
- Divide and conquer idea: use simple parts to build complex models. (e.g. multimodal densities, or piecewise-linear regressions).



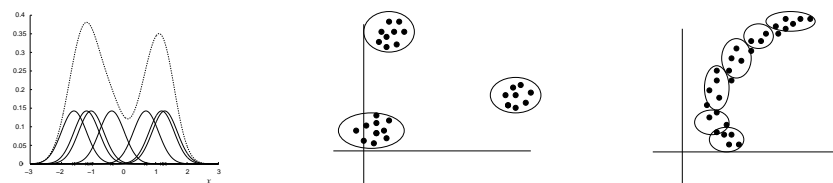
## CLUSTERING EXAMPLE: GAUSSIAN MIXTURE MODELS

- Consider a mixture of  $K$  Gaussian components:

$$p(\mathbf{x}|\theta) = \sum_k \alpha_k \mathcal{N}(\mathbf{x}|\mu_k, \Sigma_k)$$

$$p(z = k|\mathbf{x}, \theta) = \frac{\alpha_k \mathcal{N}(\mathbf{x}|\mu_k, \Sigma_k)}{\sum_j \alpha_j \mathcal{N}(\mathbf{x}|\mu_j, \Sigma_j)}$$

$$\ell(\theta; \mathcal{D}) = \sum_n \log \sum_k \alpha_k \mathcal{N}(\mathbf{x}^n|\mu_k, \Sigma_k)$$



- Density model:  $p(x|\theta)$  is a familiarity signal.  
Clustering:  $p(z|\mathbf{x}, \theta)$  is the assignment rule,  $-\ell(\theta)$  is the cost.

## MIXTURE DENSITIES

- Exactly like a classification model but the class is unobserved and so we sum it out. What we get is a perfectly valid density:

$$\begin{aligned} p(\mathbf{x}|\theta) &= \sum_{k=1}^K p(z = k|\theta_z) p(\mathbf{x}|z = k, \theta_k) \\ &= \sum_k \alpha_k p_k(\mathbf{x}|\theta_k) \end{aligned}$$

where the “mixing proportions” add to one:  $\sum_k \alpha_k = 1$ .

- We can use Bayes’ rule to compute the posterior probability of the mixture component given some data:

$$p(z = k|\mathbf{x}, \theta) = \frac{\alpha_k p_k(\mathbf{x}|\theta_k)}{\sum_j \alpha_j p_j(\mathbf{x}|\theta_j)}$$

these quantities are called *responsibilities*.

## REGRESSION EXAMPLE: MIXTURES OF EXPERTS

- Also called conditional mixtures. Exactly like a class-conditional model but the class is unobserved and so we sum it out again:

$$\begin{aligned} p(\mathbf{y}|\mathbf{x}, \theta) &= \sum_{k=1}^K p(z = k|\mathbf{x}, \theta_z) p(\mathbf{y}|z = k, \mathbf{x}, \theta_k) \\ &= \sum_k \alpha_k(\mathbf{x}|\theta_z) p_k(\mathbf{y}|\mathbf{x}, \theta_k) \end{aligned}$$

where  $\sum_k \alpha_k(\mathbf{x}) = 1 \quad \forall \mathbf{x}$ .

- Harder: must learn  $\alpha(\mathbf{x})$  (unless chose  $z$  independent of  $\mathbf{x}$ ).
- We can still use Bayes’ rule to compute the posterior probability of the mixture component given some data:

$$p(z = k|\mathbf{x}, \mathbf{y}, \theta) = \frac{\alpha_k(\mathbf{x}) p_k(\mathbf{y}|\mathbf{x}, \theta_k)}{\sum_j \alpha_j(\mathbf{x}) p_j(\mathbf{y}|\mathbf{x}, \theta_j)}$$

this function is often called the *gating function*.

## EXAMPLE: MIXTURE OF LINEAR REGRESSION EXPERTS

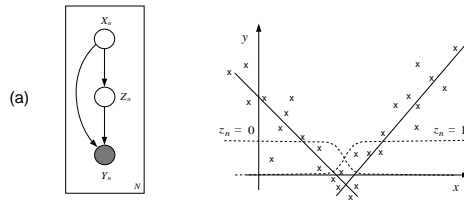
- Each expert generates data according to a linear function of the input plus additive Gaussian noise:

$$p(y|\mathbf{x}, \theta) = \sum_k \alpha_k(\mathbf{x}) \mathcal{N}(y | \beta_k^\top \mathbf{x}, \sigma_k^2)$$

- The “gate” function can be a softmax classification machine:

$$\alpha_k(\mathbf{x}) = p(z = k | \mathbf{x}) = \frac{e^{\eta_k^\top \mathbf{x}}}{\sum_j e^{\eta_j^\top \mathbf{x}}}$$

- Remember: we are not modeling the density of the inputs  $\mathbf{x}$ .



## PARAMETER CONSTRAINTS

- If we want to use general optimizations (e.g. conjugate gradient) to learn latent variable models, we often have to make sure parameters respect certain constraints. (e.g.  $\sum_k \alpha_k = 1$ ,  $\Sigma_k$  pos.definite).
- A good trick is to reparameterize these quantities in terms of unconstrained values. For mixing proportions, use the softmax:

$$\alpha_k = \frac{\exp(q_k)}{\sum_j \exp(q_j)}$$

- For covariance matrices, use the Cholesky decomposition:

$$\Sigma^{-1} = A^\top A$$

$$|\Sigma|^{-1/2} = \prod_i A_{ii}$$

where  $A$  is upper diagonal with positive diagonal:

$$A_{ii} = \exp(r_i) > 0 \quad A_{ij} = a_{ij} \quad (j > i) \quad A_{ij} = 0 \quad (j < i)$$

## GRADIENT LEARNING WITH MIXTURES

- We can learn mixture densities using gradient descent on the likelihood as usual. The gradients are quite interesting:

$$\begin{aligned} \ell(\theta) &= \log p(\mathbf{x}|\theta) = \log \sum_k \alpha_k p_k(\mathbf{x}|\theta_k) \\ \frac{\partial \ell}{\partial \theta} &= \frac{1}{p(\mathbf{x}|\theta)} \sum_k \alpha_k \frac{\partial p_k(\mathbf{x}|\theta_k)}{\partial \theta} \\ &= \sum_k \alpha_k \frac{1}{p(\mathbf{x}|\theta)} p_k(\mathbf{x}|\theta_k) \frac{\partial \log p_k(\mathbf{x}|\theta_k)}{\partial \theta} \\ &= \sum_k \alpha_k \frac{p_k(\mathbf{x}|\theta_k)}{p(\mathbf{x}|\theta)} \frac{\partial \ell_k}{\partial \theta_k} = \sum_k \alpha_k r_k \frac{\partial \ell_k}{\partial \theta_k} \end{aligned}$$

- In other words, the gradient is the *responsibility weighted sum* of the individual log likelihood gradients.

## LOGSUM

- Often you can easily compute  $b_k = \log p(\mathbf{x}|z = k, \theta_k)$ , but it will be very negative, say  $-10^6$  or smaller.
- Now, to compute  $\ell = \log p(\mathbf{x}|\theta)$  you need to compute  $\log \sum_k e^{b_k}$ . (e.g. for calculating responsibilities at test time or for learning)
- Careful! Do not compute this by doing  $\log(\text{sum}(\exp(b)))$ . You will get underflow and an incorrect answer.
- Instead do this:
  - Add a constant exponent  $B$  to all the values  $b_k$  such that the largest value comes close to the maximum exponent allowed by machine precision:  $B = \text{MAXEXPONENT} - \log(K) - \max(b)$ .
  - Compute  $\log(\text{sum}(\exp(b+B))) - B$ .
- Example: if  $\log p(x|z = 1) = -120$  and  $\log p(x|z = 2) = -120$ , what is  $\log p(x) = \log [p(x|z = 1) + p(x|z = 2)]$ ?  
Answer:  $\log[2e^{-120}] = -120 + \log 2$ .