# Corrigendum to <br> "Efficient similarity search and classification via rank aggregation" by Ronald Fagin, Ravi Kumar and D. Sivakumar, Proc. SIGMOD'03. 

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In this corrigendum, we correct an error in the paper [FKS03]. The error was discovered by Alexandr Andoni, and the corrected theorem is due to the three authors of [FKS03], along with Alexandr Andoni and Mihai Pǎtraşcu.

Theorem 4 of [FKS03] states:

> Let $D$ be a collection of $n$ points in $\mathbb{R}^{d}$. Let $r_{1}, \ldots r_{m}$ be random unit vectors in $\mathbb{R}^{d}$, where $m=\alpha \epsilon^{-2} \log n$ with a suitably chosen. Let $q \in \mathbb{R}^{d}$ be an arbitrary point, and define, for each $i$ with $1 \leq i \leq m$, the ranked list $L_{i}$ of the $n$ points in $D$ by sorting them in increasing order of their distances to the projection of $q$ along $r_{i}$. For each element $x$ of $D$, let medrank $(x)=$ median $\left(L_{1}(x), \ldots, L_{m}(x)\right)$. Let $z$ be a member of $D$ such that medrank $(x)$ is minimized. Then with probability at least $1-1 / n$, we have $d(z, q) \leq(1+\epsilon) d(x, q)$ for all $x \in D$.

As stated, the above theorem does not hold, but a slight modification of it holds. Below, we first give a counterexample to the original theorem, and then present a modified theorem.

## 1. A COUNTER-EXAMPLE

In our counter-example, we give a specific set of $n$ points in 2-dimensional space. Consider the following point set for very small $\epsilon$, illustrated in Fig. 1:

- point $q=(0,0)$, the query;
- point $p=(0,1)$, the nearest neighbor;
- point $z=(1+\epsilon, 0)$, the false nearest neighbor;
- set $H$ of 10 points $h$ all at distance $(1+\epsilon)^{2}$ from $q$, specifically $h=(1+\epsilon)^{2} \cdot\left(\frac{1}{\sqrt{2}}, \frac{1}{\sqrt{2}}\right)$;


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- set $S$ of the rest $n-12$ points, all situated at $s=$ $\left((1+\epsilon)^{2}, 0\right)$.


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Let $r$ be random unit vector in $\mathbb{R}^{2}, L$ is the list of the pointset $D$ sorted by increasing distance from $q$; and $\operatorname{rank}(x)$ is the rank of point $x$ in $L$. Then we have the following two claims.

Claim 1.1. $\operatorname{Pr}_{r}[\operatorname{rank}(z) \leq 2]>1 / 2+\Omega(\epsilon)$.
The claim follows immediately from Lemma 3 of [FKS03].
CLAIM 1.2. $\operatorname{Pr}_{r}[\operatorname{rank}(p) \geq|H|] \geq 1 / 2+1 / 16$.
It is sufficient to consider $r$ 's with non-negative $x$ coordinate, and identify $r$ 's by their angle with the $O x$ axis. First, $\operatorname{rank}(p) \leq \operatorname{rank}(z)$ if $r \in[\alpha, \beta]$, where $\alpha$ is angle formed by the perpendicular to the line connecting $q$ to midpoint of $p z$, and $\beta$ is the angle formed by the perpendicular to $p z$. We can estimate $\alpha$ and $\beta$ :

$$
\begin{gathered}
\alpha=\arctan \frac{p_{y}+z_{y}}{p_{x}+z_{x}}-\pi / 2=\arctan (1+\epsilon)-\pi / 2=-\pi / 4-\Theta(\epsilon) \\
\beta=\arctan \frac{z_{x}-p_{x}}{p_{y}-z_{y}}=\arctan (1+\epsilon)=\pi / 4+\Theta(\epsilon) .
\end{gathered}
$$

Thus, if $r \in[\alpha, \beta], \operatorname{rank}(p) \geq \operatorname{rank}(z)$, and then $\operatorname{rank}(p) \geq$ $|S|+1$.
Moreover, as we will see, if the angle of $r$ is around $-\pi / 4$, then $\operatorname{rank}(p)>\operatorname{rank}(h)$. Indeed, consider any angle $\gamma \in$ $[-\pi / 4,-\pi / 4+\pi / 16]$. Then, $|\langle p, r\rangle|=|\sin \gamma| \geq 0.5$ and $|\langle h, r\rangle|=\left|(1+\epsilon)^{2} \frac{1}{\sqrt{2}} \cdot(\sin \gamma+\cos \gamma)\right| \leq 0.2(1+\epsilon)^{2}$.

Thus, if the angle of $r$ is in the range $(-\pi / 2,-\pi / 4+$ $\pi / 16)$ or $(\beta, \pi / 2), \operatorname{rank}(p) \geq|H|$, and this happens with probability at least $\frac{\pi / 4+\pi / 16-\Theta(\epsilon)}{\pi / 2} \geq 1 / 2+1 / 16$.

Standard high concentration bounds will yield that medrank(z) $\operatorname{medrank}(p)$ with high probability. For completeness, we include one such lemma, due to Indyk:

Lemma 1.3 (cf. [Ind00], Lemma 2). Let $\mathcal{D}$ be a distribution on $\mathbb{R}$ and $F$ be its cumulative distribution function. Then, for $\epsilon, \delta>0$ and $k=O\left(\frac{\log 1 / \delta}{\epsilon^{2}}\right)$, if $X_{1} \ldots X_{k}$ are iid from $\mathcal{D}$, then $X=$ median $\left\{X_{1}, \ldots X_{k}\right\}$ satisfies $\operatorname{Pr}[F(x) \in$ $(1 / 2-\epsilon, 1 / 2+\epsilon)] \geq 1-\delta$.

## 2. A NEW ALGORITHM

To correct the theorem, we propose to use the following new function medrank $(x)$ :

$$
\operatorname{medrank}(x)=\operatorname{median}_{i}\left(\left|x r_{i}-q r_{i}\right|\right) .
$$

The resulting algorithm is presented in Fig. 2. Next, we show that this algorithm gives a $1+\epsilon$ nearest neighbor data structure.

We note that, for Step 2 of the query algorithm, we can use also other aggregation functions instead of the median function. In particular, if we use $\ell_{2}$ norm of the score vector instead of the median, then the same lemma as above holds, implied by Johnson-Lindernstrauss lemma [JL84]. Furthermore, if use $\ell_{1}$ norm of the score vector, then again the same lemma as above holds, and is implied by the $\ell_{2}$ to $\ell_{1}$ embedding of [JS82].

## 3. REFERENCES

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Preprocessing. Input: a set of points $P \subset \mathbb{R}^{d},|P|=n$, and $\epsilon>0$.

1. Choose $k=O\left(\frac{\log n}{\epsilon^{2}}\right)$ vectors $r_{i} \in \mathbb{R}^{d}, i=1 \ldots k$, where each coordinate of $r_{i}$ is drawn from a Gaussian $N(0,1)$ distribution. Vectors $r_{i}$ represent random projections.
2. Construct $k$ lists, where the $i^{\text {th }}$ list contains all the points from $p \in P$ sorted according to the value $p \cdot r_{i}$.
Query. Input: a query point $q \in \mathbb{R}^{d}$.
3. For fixed $i$ and $p \in P$, define $\operatorname{score}_{i}(p)=q \cdot r_{i}-p \cdot r_{i}$.
4. Find the point $p^{*} \in P$ that minimizes $\operatorname{median}_{i \in[k]}\left\{\left|\operatorname{score}_{i}\left(p^{*}\right)\right|\right\}$.
5. Return $p^{*}$.

Lemma 2.1. The algorithm from Figure 2 returns a $1+\epsilon$ nearest neighbor of $q$ with probability at least $1-1 / n$.

Proof. Fix some $p$ and let $\Delta=\|p-q\|_{2}$. For each $i \in[k], \operatorname{score}_{i}(p)$ is distributed as $N\left(0, \Delta^{2}\right)$, normal distribution with standard deviation $\Delta$. We will once again use Lemma 1.3 for estimating the median of iid samples.

Let $M_{p}=$ median $_{i \in[k]}\left\{\left|\operatorname{score}_{i}(p)\right|\right\}$. Applying Lemma 1.3 to the distribution $N\left(0, \Delta^{2}\right)$, we conclude that $F\left(M_{p}\right) \in$ $(1 / 2-\epsilon, 1 / 2+\epsilon)$ with probability at least $1-1 / n^{2}$, where $F$ is the cumulative of $N\left(0, \Delta^{2}\right)$. Since the value $x$ that satisfies $F(x)=1 / 2$ is $x=c \Delta$ where $c$ is an absolute constant, and $F$ has derivate $\Theta(1 / \Delta)$ around this $x$, we conclude that $M_{p} \in$ $(x-O(\epsilon \Delta), x+O(\epsilon \Delta))$. Thus, $M_{p}$ is a $1+\epsilon$ approximation to $\|q-p\|$ with probability at least $1-1 / n^{2}$.

We conclude that $M_{p}$ is a $1+\epsilon$ approximation to $\|q-p\|$ for all $p$, with probability at least $1-1 / n$. Thus the algorithm returns a $1+\epsilon$ approximate nearest neighbor with probability at least $1-1 / n$.

